

Pseudorandom Functions in Almost Constant Depth from Low-Noise LPN

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Abstract. Pseudorandom functions (PRFs) play a central role in symmetric cryptography. While in principle they can be built from any one-way functions by going through the generic HILL (SICOMP 1999) and GGM (JACM 1986) transforms, some of these steps are inherently sequential and far from practical. Naor, Reingold (FOCS 1997) and Rosen (SICOMP 2002) gave parallelizable constructions of PRFs in NC^2 and TC^0 based on concrete number-theoretic assumptions such as DDH, RSA, and factoring. Banerjee, Peikert, and Rosen (Eurocrypt 2012) constructed relatively more efficient PRFs in NC^1 and TC^0 based on “learning with errors” (LWE) for certain range of parameters. It remains an open problem whether parallelizable PRFs can be based on the “learning parity with noise” (LPN) problem for both theoretical interests and efficiency reasons (as the many modular multiplications and additions in LWE would then be simplified to AND and XOR operations under LPN).

In this paper, we give more efficient and parallelizable constructions of randomized PRFs from LPN under noise rate n^{-c} (for any constant $0 < c < 1$) and they can be implemented with a family of polynomial-size circuits with unbounded fan-in AND, OR and XOR gates of depth $\omega(1)$, where $\omega(1)$ can be any small super-constant (e.g., $\log \log \log n$ or even less). Our work complements the lower bound results by Razborov and Rudich (STOC 1994) that PRFs of beyond quasi-polynomial security are not contained in $AC^0(\text{MOD}_2)$, i.e., the class of polynomial-size, constant-depth circuit families with unbounded fan-in AND, OR, and XOR gates.

Furthermore, our constructions are security-lifting by exploiting the redundancy of low-noise LPN. We show that in addition to parallelizability (in almost constant depth) the PRF enjoys either of (or any tradeoff between) the following:

- A PRF on a weak key of sublinear entropy (or equivalently, a uniform key that leaks any $(1 - o(1))$ -fraction) has comparable security to the underlying LPN on a linear size secret.
- A PRF with key length λ can have security up to $2^{O(\lambda/\log \lambda)}$, which goes much beyond the security level of the underlying low-noise LPN. where adversary makes up to certain super-polynomial amount of queries.

1 Introduction

LEARNING PARITY WITH NOISE. The computational version of learning parity with noise (LPN) assumption with parameters $n \in \mathbb{N}$ (length of secret), $q \in \mathbb{N}$ (number of queries) and $0 < \mu < 1/2$ (noise rate) postulates that it is computationally infeasible to recover the n -bit secret $s \in \{0, 1\}^n$ given $(a \cdot s \oplus e, a)$, where a is a random $q \times n$ matrix, e follows Ber_μ^q , Ber_μ denotes the Bernoulli distribution with parameter μ (i.e., $\Pr[\text{Ber}_\mu = 1] = \mu$ and $\Pr[\text{Ber}_\mu = 0] = 1 - \mu$), ‘ \cdot ’ denotes matrix vector multiplication over $\text{GF}(2)$ and ‘ \oplus ’ denotes bitwise XOR. The decisional version of LPN simply assumes that $a \cdot s \oplus e$ is pseudorandom (i.e., computationally indistinguishable from uniform randomness) given a . The two versions are polynomially equivalent [5, 12, 36].

HARDNESS OF LPN. The computational LPN problem represents a well-known NP-complete problem “decoding random linear codes” [9] and thus its worst-case hardness is well studied. LPN was also extensively studied in learning theory, and it was shown in [24] that an efficient algorithm for LPN would allow to learn several important function classes such as 2-DNF formulas, juntas, and any function with a sparse Fourier spectrum. Under a constant noise rate (i.e., $\mu = \Theta(1)$), the best known LPN solvers [13, 40] require time and query complexity both $2^{O(n/\log n)}$. The time complexity goes up to $2^{O(n/\log \log n)}$ when restricted to $q = \text{poly}(n)$ queries [42], or even $2^{O(n)}$ given only $q = O(n)$ queries [45]. Under low noise rate $\mu = n^{-c}$ ($0 < c < 1$), the security of LPN is less well understood: on the one hand, for $q = n + O(1)$ we can already do efficient distinguishing attacks with advantage $2^{-O(n^{1-c})}$ that match the statistical distance between the LPN samples and uniform randomness (see Remark 2); on the other hand, for (even super-)polynomial q the best known attacks [7, 11, 15, 39, 54] are not asymptotically better, i.e., still at the order of $2^{\Theta(n^{1-c})}$. We mention that LPN does not succumb to known quantum algorithms, which makes it a promising candidate for “post-quantum cryptography”. Furthermore, LPN also enjoys simplicity and is more suited for weak-power devices (e.g., RFID tags) than other quantum-secure candidates such as LWE [52]¹.

LPN-BASED CRYPTOGRAPHIC APPLICATIONS. LPN was used as a basis for building lightweight authentication schemes against passive [31] and even active

¹ The inner product of LWE requires many multiplications modulo a large prime p (polynomial in the security parameter), and in contrast the same operation for LPN is simply an XOR sum of a few AND products.

adversaries [35, 36] (see [1] for a more complete literature). Recently, Kiltz et al. [38] and Dodis et al. [20] constructed randomized MACs based on the hardness of LPN, which implies a two-round authentication scheme with man-in-the-middle security. Lyubashevsky and Masny [43] gave an more efficient three-round authentication scheme from LPN (without going through the MAC transformation) and recently Cash, Kiltz, and Tessaro [16] reduced the round complexity to 2 rounds. Applebaum et al. [4] showed how to constructed a linear-stretch² pseudorandom generator (PRG) from LPN. We mention other not-so-relevant applications such as public-key encryption schemes [3, 22, 37], oblivious transfer [19], commitment schemes and zero-knowledge proofs [33], and refer to a recent survey [49] on the current state-of-the-art about LPN.

DOES LPN IMPLY LOW-DEPTH PRFS? Pseudorandom functions (PRFs) play a central role in symmetric cryptography. While in principle PRFs can be obtained via a generic transform from any one-way function [26, 29], these constructions are inherently sequential and too inefficient to compete with practical instantiations (e.g., the AES block cipher) built from scratch. Motivated by this, Naor, Reingold [46] and Rosen [47] gave direct constructions of PRFs from concrete number-theoretic assumptions (such as decision Diffie-Hellman, RSA, and factoring), which can be computed by low-depth circuits in NC^2 or even TC^0 . However, these constructions mainly established the feasibility result and are far from practical as they require extensive preprocessing and many exponentiations in large multiplicative groups. Banerjee, Peikert, and Rosen [6] constructed relatively more efficient PRFs in NC^1 and TC^0 based on the “learning with errors” (LWE) assumption. More specifically, they observed that LWE for certain range of parameters implies a deterministic variant which they call “learning with rounding” (LWR), and that LWR in turn gives rise to pseudorandom synthesizers [46], a useful tool for building low-depth PRFs. Despite that LWE is generalized from LPN, the derandomization technique used for LWE [6] does not seemingly apply to LPN, and thus it is an interesting open problem if low-depth PRFs can be based on (even a low-noise variant of) LPN (see a discussion in [49, Footnote 18]). In fact, we don’t even know how to build low-depth weak PRFs from LPN. Applebaum [4] observed that LPN implies “weak randomized pseudorandom functions”, which require independent secret coins on every function evaluation, and Akavia et al. [2] obtained weak PRFs in “ $\text{AC}^0 \circ \text{MOD}_2$ ” from a relevant non-standard hard learning assumption.

OUR CONTRIBUTIONS. In this paper, we give constructions of low-depth PRFs from low-noise LPN (see Theorem 1 below), where the noise rate n^{-c} (for any constant $0 < c < 1$) encompasses the noise level of Alekhnovich [3] (i.e., $c = 1/2$) and higher noise regime. Strictly speaking, the PRFs we obtain are not contained in $\text{AC}^0(\text{MOD}_2)^3$, but the circuit depth $\omega(1)$ can be arbitrarily small (e.g.,

² A PRG $G: \{0, 1\}^{\ell_1} \rightarrow \{0, 1\}^{\ell_2}$ has linear stretch if the stretch factor ℓ_2/ℓ_1 equals some constant greater than 1.

³ Recall that $\text{AC}^0(\text{MOD}_2)$ refers to the class of polynomial-size, constant-depth circuit families with unbounded fan-in AND, OR, and XOR gates.

$\log \log \log n$ or even less). This complements the negative result of Razborov and Rudich [51] (which is based on the works of Razborov and Smolensky [50, 53]) that PRFs with more than quasi-polynomial security do not exist in $AC^0(\text{MOD}_2)$.

Theorem 1 (main results, informal). *Assume that the LPN problem with secret length n and noise rate $\mu = n^{-c}$ (for any constant $0 < c < 1$) is $(q = 1.001n, t = 2^{O(n^{1-c})}, \epsilon = 2^{-O(n^{1-c})})$ -hard⁴. Then,*

1. *for any $d = \omega(1)$, there exists a $(q' = n^{d/3}, t - q'$ poly(n), $O(nq'\epsilon)$)-randomized-PRF on any weak key of Rényi entropy no less than $O(n^{1-c} \cdot \log n)$, or on an $n^{1-\frac{c}{2}}$ -bit uniform random key with any $(1 - \frac{O(\log n)}{n^{c/2}})$ -fraction of leakage (independent of the public coins of the PRF);*
2. *let $\lambda = \Theta(n^{1-c} \log n)$, for any $d = \omega(1)$, there exists a $(q' = \lambda^{\Theta(d)}, t' = 2^{O(\lambda/\log \lambda)}, \epsilon' = 2^{-O(\lambda/\log \lambda)})$ -randomized PRF with key length λ ;*

where both PRFs are computable by polynomial-size depth- $O(d)$ circuits with unbounded-fan-in AND, OR and XOR gates.

ON LIFTED SECURITY. Note that there is nothing special with the factor 1.001, which can be replaced with any constant greater than 1. The first parallelizable PRF has security⁵ comparable to the underlying LPN (with linear secret length) yet it uses a key of only sublinear entropy, or in the language of leakage resilient cryptography, a sublinear-size secret key with any $(1 - o(1))$ -fraction of leakage (independent of the public coins). From a different perspective, let the security parameter λ be the key length of the PRF, then the second PRF can have security up to $2^{O(\lambda/\log \lambda)}$ given any $n^{\Theta(d)}$ number of queries. We use security-preserving PRF constructions without relying on k -wise independent hash functions. This is crucial for low-depth constructions as recent works [17, 34] use (almost) $\omega(\log n)$ -wise independent hash functions, which are not known to be computable in (almost) constant-depth even with unbounded fan-in gates. We remark that circuit depth $d = \omega(1)$ is independent of the time/advantage security of PRF, and is reflected only in the query complexity $q' = n^{\Theta(d)}$. This is reasonable in many scenarios as in practice the number of queries may depend not only on adversary's computing power but also on the amount of data available for cryptanalysis. It remains open whether the dependency of query complexity on circuit depth can be fully eliminated.

BERNOULLI-LIKE RANDOMNESS EXTRACTOR/SAMPLER. Of independent interests, we propose the following randomness extractor/sampler in constant depth and they are used in the first/second PRF constructions respectively.

⁴ t and $1/\epsilon$ are upper bounded by $2^{O(n^{1-c})}$ due to known attacks.

⁵ Informally, we say that a PRF has security T if it is $1/T$ -indistinguishable from a random function for all oracle-aid distinguishers running in time T and making up to certain superpolynomial number of queries.

- A Bernoulli randomness extractor in $AC^0(MOD_2)$ that converts almost all entropy of a weak Rényi entropy source into Bernoulli noise distributions.
- A sampler in AC^0 that uses a short uniform seed and outputs a Bernoulli-like distribution of length m and noise rate μ , denoted as ψ_μ^m (see Algorithm 1).

Alekhnovich’s cryptosystem [3] considers a random distribution of length m that has exactly μm 1’s, which we denote as $\chi_{\mu m}^m$. The problem of sampling $\chi_{\mu m}^m$ dates back to [12], but the authors only mention that it can be done efficiently, and it is not known whether $\chi_{\mu m}^m$ can be sampled in $AC^0(MOD_2)$. Instead, Applebaum et al. [4] propose the following sampler for Bernoulli distribution Ber_μ^q using uniform randomness. Let $w = w_1 \cdots w_n$ be an n -bit uniform random string, and for convenience assume that μ is a negative power of 2 (i.e., $\mu = 2^{-v}$ for integer v). Let $\text{sample} : \{0, 1\}^v \rightarrow \{0, 1\}$ output the AND of its input bits, and let

$$e = (\text{sample}(w_1 \cdots w_v), \dots, \text{sample}(w_{(q-1)v+1} \cdots w_{(q-1)v+v}))$$

so that $e \sim Ber_\mu^q$ for any $q \leq \lfloor n/\log(1/\mu) \rfloor$. Note that Ber_μ has Shannon entropy $\mathbf{H}_1(Ber_\mu) = \Theta(\mu \log(1/\mu))$ (see Fact A1), and thus the above converts a $(q\mathbf{H}_1(Ber_\mu)/n) = O(\mu)$ -fraction of the entropy into Bernoulli randomness. It was observed in [4] that conditioned on e source w remains of $(1 - O(\mu))n$ bits of average min-entropy, which can be recycled into uniform randomness with a universal hash function h . That is, the two distributions are statistically close

$$(e, h(w), h) \stackrel{\approx}{\sim} (Ber_\mu^q, U_{(1-O(\mu))n}, h),$$

where U_q denotes a uniform distribution over $\{0, 1\}^q$. The work of [4] then proceeded to a construction of PRG under noise rate $\mu = \Theta(1)$. However, for $\mu = n^{-c}$ the above only samples an $O(n^{-c})$ -fraction of entropy. To convert more entropy into Bernoulli distributions, one may need to apply the above sample-then-recycle process to the uniform randomness recycled from a previous round (e.g., $h(w)$ of the first round) and repeat the process many times. However, this method is sequential and requires a circuit of depth $\Omega(n^c)$ to convert any constant fraction of entropy. We propose a more efficient and parallelizable extractor in $AC^0(MOD_2)$. As shown in Fig. 1, given any weak source of Rényi entropy $\Theta(n)$, we apply i.i.d. pairwise independent hash functions h_1, \dots, h_q (each of output length v) to w and then use sample on the bits extracted to get the Bernoulli distributions. We prove a lemma showing that this method can transform almost all entropy into Bernoulli distribution Ber_μ^q , namely, the number of extracted Bernoulli bits q can be up to $\Theta(n/\mathbf{H}_1(Ber_\mu))$. This immediately gives an equivalent formulation of the standard LPN by reusing matrix a to randomize the hash functions. For example, for each $1 \leq i \leq q$ denote by a_i the i -th row of a , let h_i be described by a_i , and let i -th LPN sample be $\langle a_i, s \rangle \oplus \text{sample}(h_i(w))$. Note that the algorithm is non-trivial as $(h_1(w), \dots, h_q(w))$ can be of length $\Theta(n^{1+c})$, which is much greater than the entropy of w .

The Bernoulli randomness extractor is used in the first PRF construction. For our second construction, we introduce a Bernoulli-like distribution ψ_μ^m that can be more efficiently sampled in AC^0 (i.e., without using XOR gates), and show that it can be used in place of Ber_μ^m with provable security.

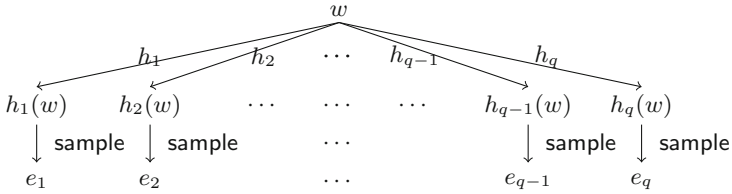


Fig. 1. An illustration of the proposed Bernoulli randomness extractor in $AC^0(\text{MOD}_2)$.

PRGS AND PRFS FROM LPN. It can be shown that standard LPN implies a variant where the secret s and noise vector e are sampled from Ber_μ^{n+q} or even ψ_μ^{n+q} . This allows us to obtain a randomized PRG G_a with short seed and polynomial stretch, where a denotes the public coin. We then use the technique of Goldreich, Goldwasser and Micali [26] with a $n^{\Theta(1)}$ -ary tree of depth $\omega(1)$ (reusing public coin a at every invocation of G_a) and construct a randomized PRF (see Definition 4) $F_{k,a}$ with input length $\omega(\log n)$, secret key k and public coin a . This already implies PRFs of arbitrary input length by Levin’s trick [41], i.e., $\bar{F}_{(k,h),a}(x) \stackrel{\text{def}}{=} F_{k,a}(h(x))$ where h is a universal hash function from any fixed-length input to $\omega(\log n)$ bits. Note that $\bar{F}_{(k,h),a}$ is computable in depth $\omega(1)$ (i.e., the depth of the GGM tree) for any small $\omega(1)$. However, the security of the above does not go beyond $n^{\omega(1)}$ due to a birthday attack. To overcome this, we use a simple and parallel method [8, 44] by running a sub-linear number of independent⁶ copies of $\bar{F}_{(k,h),a}$ and XORing their outputs, and we avoid key expansions by using pseudorandom keys (expanded using G_a or $F_{k,a}$) for all copies of $\bar{F}_{(k,h),a}$. We obtain our final security-preserving construction of PRFs by putting together all the above ingredients.

The rest of the paper is organized as follows: Sect. 2 gives background information about relevant notions and definitions. Section 3 presents the Bernoulli randomness extractor. Sections 4 and 5 give the two constructions of PRFs respectively. We include in Appendix A well-known lemmas and inequalities used, and refer to Appendix B for all the proofs omitted in the main text.

2 Preliminaries

NOTATIONS AND DEFINITIONS. We use $[n]$ to denote set $\{1, \dots, n\}$. We use capital letters⁷ (e.g., X, Y) for random variables and distributions, standard letters (e.g., x, y) for values, and calligraphic letters (e.g. \mathcal{X}, \mathcal{E}) for sets and events. The support of a random variable X , denoted by $\text{Supp}(X)$, refers to the set of values on which X takes with non-zero probability, i.e., $\{x : \Pr[X = x] > 0\}$.

⁶ By “independent” we mean that $\bar{F}_{(k,h),a}$ is evaluated on independent keys but still reusing the same public coin a .

⁷ The two exceptions are G and F , which are reserved for PRGs and PRFs respectively.

Denote by $|\mathcal{S}|$ the cardinality of set \mathcal{S} . We use Ber_μ to denote the Bernoulli distribution with parameter μ , i.e., $\Pr[\text{Ber}_\mu = 1] = \mu$, $\Pr[\text{Ber}_\mu = 0] = 1 - \mu$, while Ber_μ^q denotes the concatenation of q independent copies of Ber_μ . We use χ_i^q , $i \leq q$, to denote a uniform distribution over $\{e \in \{0, 1\}^q : |e| = i\}$, where $|e|$ denotes the Hamming weight of binary string e . For $n \in \mathbb{N}$, U_n denotes the uniform distribution over $\{0, 1\}^n$ and independent of any other random variables in consideration, and $f(U_n)$ denotes the distribution induced by applying the function f to U_n . $X \sim D$ denotes that random variable X follows distribution D . We use $s \leftarrow S$ to denote sampling an element s according to distribution S , and let $s \stackrel{\$}{\leftarrow} \mathcal{S}$ denote sampling s uniformly from set \mathcal{S} .

ENTROPY DEFINITIONS. For a random variable X and any $x \in \text{Supp}(X)$, the sample-entropy of x with respect to X is defined as

$$\mathbf{H}_X(x) \stackrel{\text{def}}{=} \log(1/\Pr[X = x])$$

from which we define the Shannon entropy, Rényi entropy and min-entropy of X respectively, i.e.,

$$\mathbf{H}_1(X) \stackrel{\text{def}}{=} \mathbb{E}_{x \leftarrow X}[\mathbf{H}_X(x)], \quad \mathbf{H}_2 \stackrel{\text{def}}{=} -\log \sum_{x \in \text{Supp}(X)} 2^{-2\mathbf{H}_X(x)}, \quad \mathbf{H}_\infty(X) \stackrel{\text{def}}{=} \min_{x \in \text{Supp}(X)} \mathbf{H}_X(x).$$

For $0 < \mu < 1/2$, let $\mathbf{H}(\mu) \stackrel{\text{def}}{=} \mu \log(1/\mu) + (1 - \mu) \log(1/(1 - \mu))$ be the binary entropy function so that $\mathbf{H}(\mu) = \mathbf{H}_1(\text{Ber}_\mu)$. We know that $\mathbf{H}_1(X) \geq \mathbf{H}_2(X) \geq \mathbf{H}_\infty(X)$ with equality when X is uniformly distributed. A random variable X of length n is called an (n, λ) -Rényi entropy (resp., min-entropy) source if $\mathbf{H}_2(X) \geq \lambda$ (resp., $\mathbf{H}_\infty(X) \geq \lambda$). The *statistical distance* between X and Y , denoted by $\text{SD}(X, Y)$, is defined by

$$\text{SD}(X, Y) \stackrel{\text{def}}{=} \frac{1}{2} \sum_x |\Pr[X = x] - \Pr[Y = x]|$$

We use $\text{SD}(X, Y|Z)$ as a shorthand for $\text{SD}((X, Z), (Y, Z))$.

SIMPLIFYING NOTATIONS. To simplify the presentation, we use the following simplified notations. Throughout, n is the security parameter and most other parameters are functions of n , and we often omit n when clear from the context. For example, $\mu = \mu(n) \in (0, 1/2)$, $q = q(n) \in \mathbb{N}$, $t = t(n) > 0$, $\epsilon = \epsilon(n) \in (0, 1)$, and $m = m(n) = \text{poly}(n)$, where poly refers to some polynomial.

Definition 1 (Computational/decisional LPN). *Let n be a security parameter, and let μ, q, t and ϵ all be functions of n . The **decisional LPN** $_{\mu, n}$ problem (with secret length n and noise rate μ) is (q, t, ϵ) -hard if for every probabilistic distinguisher D running in time t we have*

$$\left| \Pr_{A, S, E} [D(A, A \cdot S \oplus E) = 1] - \Pr_{A, U_q} [D(A, U_q) = 1] \right| \leq \epsilon \quad (1)$$

where $A \sim U_{qn}$ is a $q \times n$ matrix, $S \sim U_n$ and $E \sim \text{Ber}_\mu^q$. The **computational LPN $_{\mu,n}$** problem is (q, t, ϵ) -hard if for every probabilistic algorithm D running in time t we have

$$\Pr_{A,S,E} [D(A, A \cdot S \oplus E) = (S, E)] \leq \epsilon,$$

where $A \sim U_{qn}$, $S \sim U_n$ and $E \sim \text{Ber}_\mu^q$.

Definition 2 (LPN variants). The decisional/computational **X-LPN $_{\mu,n}$** is defined as per Definition 1 accordingly except that (S, E) follows distribution X . Note that standard LPN $_{\mu,n}$ is a special case of X-LPN $_{\mu,n}$ for $X \sim (U_n, \text{Ber}_\mu^q)$.

In respect of the randomized feature of LPN, we generalize standard PRGs/PRFs to equivalent randomized variants, where the generator/function additionally uses some public coins for randomization, and that seed/key can be sampled from a weak source (independent of the public coins).

Definition 3 (Randomized PRGs on weak seeds). Let $\lambda \leq \ell_1 < \ell_2, \ell_3, t, \epsilon$ be functions of security parameter n . An efficient function family ensemble $\mathcal{G} = \{G_a : \{0, 1\}^{\ell_1} \rightarrow \{0, 1\}^{\ell_2}, a \in \{0, 1\}^{\ell_3}\}_{n \in \mathbb{N}}$ is a (t, ϵ) randomized PRG on (ℓ_1, λ) -weak seed if for every probabilistic distinguisher D of running time t and every (ℓ_1, λ) -Rényi entropy source K it holds that

$$\left| \Pr_{K, A \sim U_{\ell_3}} [D(G_A(K), A) = 1] - \Pr_{U_{\ell_2}, A \sim U_{\ell_3}} [D(U_{\ell_2}, A) = 1] \right| \leq \epsilon.$$

The stretch factor of \mathcal{G} is ℓ_2/ℓ_1 . Standard (deterministic) PRGs are implied by defining $G'(k, a) \stackrel{\text{def}}{=} (G_a(k), a)$ for a uniform random k .

Definition 4 (Randomized PRFs on weak keys). Let $\lambda \leq \ell_1, \ell_2, \ell_3, \ell, t, \epsilon$ be functions of security parameter n . An efficient function family ensemble $\mathcal{F} = \{F_{k,a} : \{0, 1\}^\ell \rightarrow \{0, 1\}^{\ell_2}, k \in \{0, 1\}^{\ell_1}, a \in \{0, 1\}^{\ell_3}\}_{n \in \mathbb{N}}$ is a (q, t, ϵ) randomized PRF on (ℓ_1, λ) -weak key if for every oracle-aided probabilistic distinguisher D of running time t and bounded by q queries and for every (ℓ_1, λ) -Rényi entropy source K we have

$$\left| \Pr_{K, A \sim U_{\ell_3}} [D^{F_{K,A}}(A) = 1] - \Pr_{R, A \sim U_{\ell_3}} [D^R(A) = 1] \right| \leq \epsilon(n),$$

where R denotes a random function distribution ensemble mapping from ℓ bits to ℓ_2 bits. Standard PRFs are a special case for empty a (or keeping $k' = (k, a)$ secret) on uniformly random key.

Definition 5 (Universal hashing). A function family $\mathcal{H} = \{h_a : \{0, 1\}^n \rightarrow \{0, 1\}^m, a \in \{0, 1\}^l\}$ is **universal** if for any $x_1 \neq x_2 \in \{0, 1\}^n$ it holds that

$$\Pr_{a \stackrel{\$}{\leftarrow} \{0, 1\}^l} [h_a(x_1) = h_a(x_2)] \leq 2^{-m}.$$

Definition 6 (Pairwise independent hashing). A function family $\mathcal{H} = \{h_a : \{0, 1\}^n \rightarrow \{0, 1\}^m, a \in \{0, 1\}^l\}$ is **pairwise independent** if for any $x_1 \neq x_2 \in \{0, 1\}^n$ and any $v \in \{0, 1\}^{2m}$ it holds that

$$\Pr_{a \leftarrow \{0, 1\}^l} [(h_a(x_1), h_a(x_2)) = v] = 2^{-2m}.$$

CONCRETE CONSTRUCTIONS. We know that for every $m \leq n$ there exists a pairwise independent (and universal) \mathcal{H} with description length $l = \Theta(n)$, where every $h \in \mathcal{H}$ can be computed in $\text{AC}^0(\text{MOD}_2)$. For example, \mathcal{H}_1 and \mathcal{H}_2 defined below are universal and pairwise independent respectively:

$$\mathcal{H}_1 = \{h_a : \{0, 1\}^n \rightarrow \{0, 1\}^m \mid h_a(x) \stackrel{\text{def}}{=} a \cdot x, a \in \{0, 1\}^{n+m-1}\}$$

$$\mathcal{H}_2 = \{h_{a,b} : \{0, 1\}^n \rightarrow \{0, 1\}^m \mid h_{a,b}(x) \stackrel{\text{def}}{=} a \cdot x \oplus b, a \in \{0, 1\}^{n+m-1}, b \in \{0, 1\}^m\}$$

where $a \in \{0, 1\}^{n+m-1}$ is interpreted as an $m \times n$ Toeplitz matrix and ‘ \cdot ’ and ‘ \oplus ’ denote matrix-vector multiplication and addition over $\text{GF}(2)$ respectively.

3 Bernoulli Randomness Extraction in $\text{AC}^0(\text{MOD}_2)$

First, we state below a variant of the lemma (e.g., [28]) that taking sufficiently many samples of i.i.d. random variables yields an “almost flat” joint random variable, i.e., the sample-entropy of most values is close to the Shannon entropy of the joint random variable. The proof is included in Appendix B for completeness.

Lemma 1 (Flattening Shannon entropy). For any $n \in \mathbb{N}$, $0 < \mu < 1/2$ and for any $\Delta > 0$ define

$$\mathcal{E} \stackrel{\text{def}}{=} \{e \in \{0, 1\}^q : \mathbf{H}_{\text{Ber}_\mu^q}(e) \leq (1 + \Delta)q\mathbf{H}(\mu)\}. \quad (2)$$

Then, we have $\Pr[\text{Ber}_\mu^q \in \mathcal{E}] \geq 1 - \exp^{-\frac{\min(\Delta, \Delta^2)\mu q}{3}}$.

Lemma 2 states that the proposed Bernoulli randomness extractor (see Fig. 1) extracts almost all entropy from a Rényi entropy (or min-entropy) source. We mention that the extractor can be considered as a parallelized version of the random bits recycler of Impagliazzo and Zuckerman [32] and the proof technique is also closely relevant to the crooked leftover hash lemma [14, 21].

Lemma 2 (Bernoulli randomness extraction). For any $m, v \in \mathbb{N}$ and $0 < \mu \leq 1/2$, let $W \in \mathcal{W}$ be any $(\lceil \log |\mathcal{W}| \rceil, m)$ -Rényi entropy source, let \mathcal{H} be a family of pairwise independent hash functions mapping from \mathcal{W} to $\{0, 1\}^v$, let $\mathbf{H} = (H_1, \dots, H_q)$ be a vector of i.i.d. random variables such that each H_i is uniformly distributed over \mathcal{H} , let $\text{sample} : \{0, 1\}^v \rightarrow \{0, 1\}$ be any Boolean function such that $\text{sample}(U_v) \sim \text{Ber}_\mu$. Then, for any constant $0 < \Delta \leq 1$ it holds that

$$\text{SD}(\text{Ber}_\mu^q, \text{sample}(\mathbf{H}(W)) \mid \mathbf{H}) \leq 2^{\binom{(1+\Delta)q\mathbf{H}(\mu)-m}{2}} + \exp^{-\frac{\Delta^2\mu q}{3}},$$

where

$$\text{sample}(\mathbf{H}(W)) \stackrel{\text{def}}{=} (\text{sample}(H_1(W)), \dots, \text{sample}(H_q(W))).$$

Remark 1 (On entropy loss). The amount of entropy extracted (i.e., $q\mathbf{H}(\mu)$) can be almost as large as entropy of the source (i.e., m) by setting $m = (1+2\Delta)q\mathbf{H}(\mu)$ for any arbitrarily small constant Δ . Further, the leftover hash lemma falls into a special case for $v = 1$ (sample being an identity function) and $\mu = 1/2$.

Proof. Let set \mathcal{E} be defined as in (2). For any $e \in \{0, 1\}^q$ and $\mathbf{h} \in \mathcal{H}^q$, use short-hands $p_{\mathbf{h}} \stackrel{\text{def}}{=} \Pr[\mathbf{H} = \mathbf{h}]$, $p_{e|\mathbf{h}} \stackrel{\text{def}}{=} \Pr[\text{sample}(\mathbf{h}(W)) = e]$ and $p_e \stackrel{\text{def}}{=} \Pr[\text{Ber}_{\mu}^q = e]$. We have

$$\begin{aligned} & \text{SD}((\text{Ber}_{\mu}^q, \mathbf{H}), (\text{sample}(\mathbf{H}(W)), \mathbf{H})) \\ &= \frac{1}{2} \sum_{\mathbf{h} \in \mathcal{H}^q, e \in \mathcal{E}} p_{\mathbf{h}} |p_{e|\mathbf{h}} - p_e| + \frac{1}{2} \sum_{\mathbf{h} \in \mathcal{H}^q, e \notin \mathcal{E}} p_{\mathbf{h}} |p_{e|\mathbf{h}} - p_e| \\ &\leq \frac{1}{2} \sum_{\mathbf{h} \in \mathcal{H}^q, e \in \mathcal{E}} (\sqrt{p_{\mathbf{h}} \cdot p_e}) \cdot \left(\sqrt{\frac{p_{\mathbf{h}}}{p_e}} |p_{e|\mathbf{h}} - p_e| \right) \\ &\quad + \frac{1}{2} \left(\sum_{\mathbf{h} \in \mathcal{H}^q, e \notin \mathcal{E}} p_{\mathbf{h}} p_{e|\mathbf{h}} + \sum_{\mathbf{h} \in \mathcal{H}^q, e \notin \mathcal{E}} p_{\mathbf{h}} p_e \right) \\ &\leq \frac{1}{2} \sqrt{\left(\sum_{\mathbf{h} \in \mathcal{H}^q, e \in \mathcal{E}} p_{\mathbf{h}} \cdot p_e \right) \cdot \left(\sum_{\mathbf{h} \in \mathcal{H}^q, e \in \mathcal{E}} \frac{p_{\mathbf{h}}}{p_e} \cdot (p_{e|\mathbf{h}} - p_e)^2 \right)} + \Pr[\text{Ber}_{\mu}^q \notin \mathcal{E}] \\ &\leq \frac{1}{2} \sqrt{1 \cdot \sum_{e \in \mathcal{E}} \left(\sum_{\mathbf{h} \in \mathcal{H}^q} \frac{p_{\mathbf{h}} p_{e|\mathbf{h}}^2}{p_e} - 2 \sum_{\mathbf{h} \in \mathcal{H}^q} p_{\mathbf{h}} p_{e|\mathbf{h}} + \sum_{\mathbf{h} \in \mathcal{H}^q} p_{\mathbf{h}} p_e \right)} + \exp^{-\frac{\Delta^2 \mu q}{3}} \\ &\leq \frac{1}{2} \sqrt{|\mathcal{E}| \cdot 2^{-m}} + \exp^{-\frac{\Delta^2 \mu q}{3}} \\ &\leq 2^{\frac{(1+\Delta)q\mathbf{H}(\mu)-m}{2}} + \exp^{-\frac{\Delta^2 \mu q}{3}}, \end{aligned}$$

where the second inequality is Cauchy-Schwarz, i.e., $|\sum a_i b_i| \leq \sqrt{(\sum a_i^2) \cdot (\sum b_i^2)}$ and (3) below, the third inequality follows from Lemma 1, and the fourth inequality is due to (4) and (5), i.e., fix any e (and thus fix p_e as well) we can substitute $p_e \cdot (2^{-m} + p_e)$ for $\sum_{\mathbf{h} \in \mathcal{H}^q} p_{\mathbf{h}} p_{e|\mathbf{h}}^2$, and p_e for both $\sum_{\mathbf{h} \in \mathcal{H}^q} p_{\mathbf{h}} p_{e|\mathbf{h}}$ and $\sum_{\mathbf{h} \in \mathcal{H}^q} p_{\mathbf{h}} p_e$, and the last inequality follows from the definition of \mathcal{E} (see (2))

$$|\mathcal{E}| \leq 1/\min_{e \in \mathcal{E}} \Pr[\text{Ber}_{\mu}^q = e] \leq 2^{(1+\Delta)q\mathbf{H}(\mu)}$$

which completes the proof.

Claim 1

$$\sum_{\mathbf{h} \in \mathcal{H}^q, e \notin \mathcal{E}} p_{\mathbf{h}} p_{e|\mathbf{h}} = \sum_{\mathbf{h} \in \mathcal{H}^q, e \notin \mathcal{E}} p_{\mathbf{h}} p_e = \Pr[\text{Ber}_{\mu}^q \notin \mathcal{E}] \tag{3}$$

$$\forall \mathbf{e} \in \{0, 1\}^q : \sum_{h \in \mathcal{H}^q} p_h p_{\mathbf{e}|h}^2 \leq p_{\mathbf{e}} \cdot (2^{-m} + p_{\mathbf{e}}) \quad (4)$$

$$\forall \mathbf{e} \in \{0, 1\}^q : \sum_{h \in \mathcal{H}^q} p_h p_{\mathbf{e}|h} = \sum_{h \in \mathcal{H}^q} p_h p_{\mathbf{e}} = p_{\mathbf{e}} \quad (5)$$

Proof. Let $\mathbf{H}(W) \stackrel{\text{def}}{=} (H_1(W), \dots, H_q(W))$. The pairwise independence of \mathcal{H} implies that

$$\mathbf{H}(W) \sim (U_v^1, \dots, U_v^q)$$

holds even conditioned on any fixing of $W = w$, and thus $\text{sample}(\mathbf{H}(W)) \sim \text{Ber}_{\mu}^q$. We have

$$\sum_{h \in \mathcal{H}^q, \mathbf{e} \notin \mathcal{E}} p_h p_{\mathbf{e}|h} = \Pr[\text{sample}(\mathbf{H}(W)) \notin \mathcal{E}] = \Pr[\text{Ber}_{\mu}^q \notin \mathcal{E}],$$

$$\forall \mathbf{e} \in \{0, 1\}^q : \sum_{h \in \mathcal{H}^q} p_h p_{\mathbf{e}|h} = \Pr[\text{sample}(\mathbf{H}(W)) = \mathbf{e}] = \Pr[\text{Ber}_{\mu}^q = \mathbf{e}] = p_{\mathbf{e}},$$

$$\sum_{h \in \mathcal{H}^q, \mathbf{e} \notin \mathcal{E}} p_h p_{\mathbf{e}} = \sum_{h \in \mathcal{H}^q} p_h \cdot \sum_{\mathbf{e} \notin \mathcal{E}} p_{\mathbf{e}} = \Pr[\text{Ber}_{\mu}^q \notin \mathcal{E}],$$

$$\forall \mathbf{e} \in \{0, 1\}^q : \sum_{h \in \mathcal{H}^q} p_h p_{\mathbf{e}} = p_{\mathbf{e}} \cdot \sum_{h \in \mathcal{H}^q} p_h = p_{\mathbf{e}}.$$

Now fix any $\mathbf{e} \in \{0, 1\}^q$, and let W_1 and W_2 be random variables that are i.i.d. to W , we have

$$\begin{aligned} & \sum_{h \in \mathcal{H}^q} p_h p_{\mathbf{e}|h}^2 \\ &= \Pr_{W_1, W_2, \mathbf{H}}[\text{sample}(\mathbf{H}(W_1)) = \text{sample}(\mathbf{H}(W_2)) = \mathbf{e}] \\ &\leq \Pr_{W_1, W_2}[W_1 = W_2] \cdot \Pr_{W_1, \mathbf{H}}[\text{sample}(\mathbf{H}(W_1)) = \mathbf{e}] \\ &\quad + \Pr_{\mathbf{H}}[\text{sample}(\mathbf{H}(w_1)) = \text{sample}(\mathbf{H}(w_2)) = \mathbf{e} \mid w_1 \neq w_2] \\ &\leq 2^{-m} \cdot p_{\mathbf{e}} + \Pr[\text{Ber}_{\mu}^q = \mathbf{e}]^2 = 2^{-m} \cdot p_{\mathbf{e}} + p_{\mathbf{e}}^2, \end{aligned}$$

where the second inequality is again due to the pairwise independence of \mathcal{H} , i.e., for any $w_1 \neq w_2$, $\mathbf{H}(w_1)$ and $\mathbf{H}(w_2)$ are i.i.d. to (U_v^1, \dots, U_v^q) and thus the two distributions $\text{sample}(\mathbf{H}(w_1))$ and $\text{sample}(\mathbf{H}(w_2))$ are i.i.d. to Ber_{μ}^q .

4 Parallelizable PRFs on Weak Keys

4.1 A Succinct Formulation of LPN

The authors of [22] observed that the secret of LPN is not necessary to be uniformly random and can be replaced with a Bernoulli distribution. We state a more quantitative version (than [22, Problem 2]) in Lemma 3 that $\text{Ber}_{\mu}^{m+q}\text{-LPN}_{\mu, n}$ (see Definition 2) is implied by standard LPN for nearly the same parameters except that standard LPN needs n more samples. The proof follows by a simple reduction and is included in Appendix B.

Lemma 3. *Assume that the decisional (resp., computational) LPN $_{\mu,n}$ problem is (q, t, ϵ) -hard, then the decisional (resp., computational) Ber_{μ}^{n+q} -LPN $_{\mu,n}$ problem is at least $(q - (n + 2), t - \text{poly}(n + q), 2\epsilon)$ -hard.*

Remark 2 (On the security of low-noise LPN). For $\mu = n^{-c}$, a trivial statistical test suggests (by the piling-up lemma) that any single sample of decisional Ber_{μ}^{n+q} -LPN $_{\mu,n}$ is $(1/2 + 2^{-O(n^{1-c})})$ -biased to 0. In other words, decisional Ber_{μ}^{n+q} -LPN $_{\mu,n}$ is no more than $(q = 1, t = O(1), \epsilon = 2^{-O(n^{1-c})})$ -hard and thus it follows (via the reduction of Lemma 3) that decisional LPN $_{\mu,n}$ cannot have indistinguishability beyond $(q = n + 3, t = \text{poly}(n), \epsilon = 2^{-O(n^{1-c})})$. Asymptotically, this is also the current state-of-the-art attack on low-noise LPN using $q = \text{poly}(n)$ or even more samples.

4.2 A Direct Construction in Almost Constant Depth

To build a randomized PRG (on weak source w) from the succinct LPN, we first sample Bernoulli vector (s, e) from w (using random coins a), and then output $a \cdot s \oplus e$. Theorem 2 states that the above yields a randomized PRG on weak seed w and public coin a .

Theorem 2 (randomized PRGs from LPN). *Let n be a security parameter, let $\delta > 0$ be any constant, and let $\mu = n^{-c}$ for any $0 < c < 1$. Assume that decisional LPN $_{\mu,n}$ problem is $((1 + 2\delta)n, t, \epsilon)$ -hard, then $\mathcal{G} = \{G_a : \{0, 1\}^{n^{1-\frac{\delta}{2}}} \rightarrow \{0, 1\}^{\delta n}, a \in \{0, 1\}^{\delta n \times n}\}_{n \in \mathbb{N}}$, where*

$$G_a(w) = a \cdot s \oplus e, s \in \{0, 1\}^n, e \in \{0, 1\}^{\delta n}$$

and $(s, e) = \text{sample}(\mathbf{h}_a(w))$, is a $(t - \text{poly}(n), O(\epsilon))$ -randomized PRG on $(n^{1-\frac{\delta}{2}}, 4c(1 + \delta^2)n^{1-c} \cdot \log n)$ -weak seed with stretch factor $\delta \cdot n^{\frac{\delta}{2}}$.

Proof. We have by Lemma 3 that $((1 + 2\delta)n, t, \epsilon)$ -hard decisional LPN $_{\mu,n}$ implies $(\delta n, t - \text{poly}(n), 2\epsilon)$ -hard decisional $\text{Ber}_{\mu}^{n+\delta n}$ -LPN $_{\mu,n}$, so the conclusion follows if we could sample $(s, e) \stackrel{\$}{\leftarrow} \text{Ber}_{\mu}^{n+\delta n}$ from w . This follows from Lemma 2 by choosing $q = n + \delta n$, $\Delta = \delta$, and $m = 4c(1 + \delta)^2 n^{1-c} \cdot \log n$ such that the sampled noise vector is statistically close to $\text{Ber}_{\mu}^{n+\delta n}$ except for an error bounded by

$$\begin{aligned} & 2^{((1+\Delta)q\mathbf{H}(\mu)-m)/2} + \exp^{-\frac{\Delta^2\mu q}{3}} \\ & \leq 2^{((1+\delta)^2 n\mathbf{H}(\mu)-2(1+\delta)^2 n\mathbf{H}(\mu))/2} + 2^{-\Omega(n^{1-c})} \\ & = 2^{-\Omega(n^{1-c} \cdot \log n)} + 2^{-\Omega(n^{1-c})} \\ & = 2^{-\Omega(n^{1-c})} \end{aligned}$$

where recall by Fact A1 that $\mu \log(1/\mu) < \mathbf{H}(\mu) < \mu(\log(1/\mu) + 2)$ and thus $m > 2(1 + \delta^2)n^{1-c}(c \log n + 2) > 2(1 + \delta^2)n\mathbf{H}(\mu)$. We omit the above term since $\epsilon = 2^{-O(n^{1-c})}$ (see Remark 2).

We state a variant of the theorem by Goldreich, Goldwasser and Micali [26] on building PRFs from PRGs, where we consider PRGs with stretch factor 2^v for $v = O(\log n)$ (i.e., a balanced 2^v -ary tree) and use randomized (instead of deterministic) PRG G_a , reusing public coin a at every invocation of G_a .

Theorem 3 (PRFs from PRGs [26]). *Let n be a security parameter, let $v = O(\log n)$, $\lambda \leq m = n^{O(1)}$, $\lambda = \text{poly}(n)$, $t = t(n)$ and $\epsilon = \epsilon(n)$. Let $\mathcal{G} = \{G_a : \{0, 1\}^m \rightarrow \{0, 1\}^{2^v \cdot m}, a \in \mathcal{A}\}_{n \in \mathbb{N}}$ be a (t, ϵ) randomized PRG (with stretch factor 2^v) on (m, λ) -weak seed. Parse $G_a(k)$ as 2^v blocks of m -bit strings:*

$$G_a(k) \stackrel{\text{def}}{=} G_a^{0 \dots 00}(k) \| G_a^{0 \dots 01}(k) \| \dots \| G_a^{1 \dots 11}(k)$$

where $G_a^{i_1 \dots i_v}(k)$ denotes the $(i_1 \dots i_v)$ -th m -bit block of $G_a(k)$. Then, for any $d \leq \text{poly}(n)$ and $q = q(n)$, the function family ensemble $\mathcal{F} = \{F_{k,a} : \{0, 1\}^{dv} \rightarrow \{0, 1\}^{2^v \cdot m}, k \in \{0, 1\}^m, a \in \mathcal{A}\}_{n \in \mathbb{N}}$, where

$$F_{k,a}(x_1 \dots x_{dv}) \stackrel{\text{def}}{=} G_a(G_a^{x_{(d-1)v+1} \dots x_{dv}}(\dots G_a^{x_{v+1} \dots x_{2v}}(G_a^{x_1 \dots x_v}(k)) \dots)),$$

is a $(q, t - q \cdot \text{poly}(n), dq\epsilon)$ randomized PRF on (m, λ) -weak key.

ON POLYNOMIAL-SIZE CIRCUITS. The above GGM tree has $\Theta(2^{dv})$ nodes and thus it may seem that for $dv = \omega(\log n)$ we need a circuit of super-polynomial size to evaluate $F_{k,p}$. This is not necessary since we can represent the PRF in the following alternative form:

$$F_{k,a} = G_a \circ \underbrace{\text{mux}_{x_{(d-1)v+1} \dots x_{dv}} \circ G_a}_{G_a^{x_{(d-1)v+1} \dots x_{dv}}} \circ \dots \circ \underbrace{\text{mux}_{x_{v+1} \dots x_{2v}} \circ G_a}_{G_a^{x_{v+1} \dots x_{2v}}} \circ \underbrace{\text{mux}_{x_1 \dots x_v} \circ G_a}_{G_a^{x_1 \dots x_v}}$$

where ‘ \circ ’ denotes function composition, each multiplexer $\text{mux}_{i_1 \dots i_v} : \{0, 1\}^{2^v m} \rightarrow \{0, 1\}^m$ simply selects as output the $(i_1 \dots i_v)$ -th m -bit block of its input, and it can be implemented with $O(2^v \cdot m) = \text{poly}(n)$ NOT and (unbounded fan-in) AND/OR gates of constant depth. Thus, for $v = O(\log n)$ function $F_{k,p}$ can be evaluated with a polynomial-size circuit of depth $O(d)$.

Lemma 4 (Levin’s trick [41]). *For any $\ell \leq n \in \mathbb{N}$, let R_1 be a random function distribution over $\{0, 1\}^\ell \rightarrow \{0, 1\}^n$, let \mathcal{H} be a family of universal hash functions from n bits to ℓ bits, and let H_1 be a function distribution uniform over \mathcal{H} . Let $R_1 \circ H_1(x) \stackrel{\text{def}}{=} R_1(H_1(x))$ be a function distribution over $\{0, 1\}^n \rightarrow \{0, 1\}^n$. Then, for any $q \in \mathbb{N}$ and any oracle aided D bounded by q queries, we have*

$$\left| \Pr_{R_1, H_1} [D^{R_1 \circ H_1} = 1] - \Pr_R [D^R = 1] \right| \leq \frac{q^2}{2^{\ell+1}},$$

where R is a random function distribution from n bits to n bits.

Theorem 4 (A direct PRF). *Let n be a security parameter, and let $\mu = n^{-c}$ for constant $0 < c < 1$. Assume that decisional $\text{LPN}_{\mu, n}$ problem is $(\alpha n, t, \epsilon)$ -hard*

for any constant $\alpha > 1$, then for any (efficiently computable) $d = \omega(1) \leq O(n)$ and any $q \leq n^{d/3}$ there exists a $(q, t - q \text{poly}(n), O(dq\epsilon) + q^2n^{-d})$ -randomized PRF on $(n^{1-\frac{\epsilon}{2}}, O(n^{1-c} \log n))$ ⁸-weak key

$$\bar{\mathcal{F}} = \{\bar{F}_{k,a} : \{0, 1\}^n \rightarrow \{0, 1\}^n, k \in \{0, 1\}^{n^{1-\frac{\epsilon}{2}}}, a \in \{0, 1\}^{O(n^2)}\}_{n \in \mathbb{N}} \quad (6)$$

which is computable by a uniform family of polynomial-size depth- $O(d)$ circuits with unbounded-fan-in AND, OR and XOR gates.

Proof. For $\mu = n^{-c}$, we have by Theorem 2 that the decisional $(\alpha n, t, \epsilon)$ -hard LPN $_{\mu, n}$ implies a $(t - \text{poly}(n), O(\epsilon))$ randomized PRG in $\text{AC}^0(\text{MOD}_2)$ on $(n^{1-\frac{\epsilon}{2}}, O(n^{1-c} \log n))$ -weak seed k and public coin $a \in \{0, 1\}^{O(n^2)}$ with stretch factor $2^v = n^{\frac{\epsilon}{2}}$. We plug it into the GGM construction (see Theorem 3) with tree depth $d' = 2d/c$ to get a $(q, t - q \text{poly}(n), O(dq\epsilon))$ randomized PRF on weak keys (of same parameters) with input length $d'v = d \log n$ and output length $2^v \cdot n^{1-\frac{\epsilon}{2}} = n$ as below:

$$\mathcal{F} = \{F_{k,a} : \{0, 1\}^{d \log n} \rightarrow \{0, 1\}^n, k \in \{0, 1\}^{n^{1-\frac{\epsilon}{2}}}, a \in \{0, 1\}^{O(n^2)}\}_{n \in \mathbb{N}}. \quad (7)$$

Now we expand k (e.g., by evaluating $F_{k,a}$ on a few fixed points) into a pseudo-random (\bar{k}, \bar{h}_1) , where $\bar{k} \in \{0, 1\}^{n^{1-\frac{\epsilon}{2}}}$ and \bar{h}_1 describes a universal hash function from n bits to $\ell = d \log n$ bits. Motivated by Levin's trick, we define a domain-extended PRF $\bar{F}_{k,a}(x) \stackrel{\text{def}}{=} F_{\bar{k},a} \circ \bar{h}_1(x)$. For any oracle-aided distinguisher D running in time $t - q \text{poly}(n)$ and making q queries, denote with $\delta_D(F_1, F_2) \stackrel{\text{def}}{=} |\Pr[D^{F_1}(A) = 1] - \Pr[D^{F_2}(A) = 1]|$ the advantage of D (who gets public coin A as additional input) in distinguishing between function oracles F_1 and F_2 . Therefore, we have by a triangle inequality

$$\begin{aligned} \delta_D(F_{\bar{K},A} \circ \bar{H}_1, R) &\leq \delta_D(F_{\bar{K},A} \circ \bar{H}_1, F_{K,A} \circ H_1) + \delta_D(F_{K,A} \circ H_1, R_1 \circ H_1) \\ &\quad + \delta_D(R_1 \circ H_1, R) \\ &\leq O(dq\epsilon) + q^2n^{-d}, \end{aligned}$$

where advantage is upper bounded by three terms, namely, the indistinguishability between (\bar{K}, \bar{H}_1) and truly random (K, H_1) , that between $F_{K,A}$ and random function R_1 (of the same input/output lengths as $F_{K,A}$), and that due to Lemma 4. Note that A is independent of R_1, H_1 and R .

4.3 Going Beyond the Birthday Barrier

Unfortunately, for small $d = \omega(1)$ the security of the above PRF does not go beyond super-polynomial (cf. term q^2n^{-d}) due to a birthday attack. This situation can be handled using security-preserving constructions. Note the techniques from [17, 34] need (almost) $\Omega(d \log n)$ -wise independent hash functions which we don't know how to compute with unbounded fan-in gates of depth $O(d)$. Thus,

⁸ Here the big-Oh omits a constant dependent on c and α .

we use a more intuitive and depth-preserving approach below by simply running a few independent copies and XORing their outputs. The essential idea dates back to [8, 44] and the technique receives renewed interest recently in some different contexts [23, 25]. We mention that an alternative (and possibly more efficient) approach is to use the second security-preserving domain extension technique from [10] that requires a few pairwise independent hash functions and makes only a constant number of calls to the underlying small-domain PRFs. This yields the PRF stated in Theorem 5.

Lemma 5 (Generalized Levin’s Trick [8, 44]). *For any $\kappa, \ell \leq n \in \mathbb{N}$, let R_1, \dots, R_κ be independent random function distributions over $\{0, 1\}^\ell \rightarrow \{0, 1\}^n$, let \mathcal{H} be a family of universal hash functions from n bits to ℓ bits, and let H_1, \dots, H_κ be independent function distributions all uniform over \mathcal{H} . Let $F_{\mathbf{R}, \mathbf{H}}$ be a function distribution (induced by $\mathbf{R} = (R_1, \dots, R_\kappa)$ and $\mathbf{H} = (H_1, \dots, H_\kappa)$) over $\{0, 1\}^n \rightarrow \{0, 1\}^n$ defined as*

$$F_{\mathbf{R}, \mathbf{H}}(x) \stackrel{\text{def}}{=} \bigoplus_{i=1}^{\kappa} R_i(H_i(x)). \tag{8}$$

Then, for any $q \in \mathbb{N}$ and any oracle aided D bounded by q queries, we have

$$|\Pr[D^{F_{\mathbf{R}, \mathbf{H}}} = 1] - \Pr[D^R = 1]| \leq \frac{q^{\kappa+1}}{2^{\kappa\ell}}$$

where R is a random function distribution over $\{0, 1\}^n \rightarrow \{0, 1\}^n$.

Finally, we get the first security-preserving construction below. To have comparable security to LPN with secret size n , it suffices to use a key of entropy $O(n^{1-c} \cdot \log n)$, or a uniform key of size $n^{1-\frac{c}{2}}$ with any $(1 - O(n^{-\frac{c}{2}} \log n))$ -fraction of leakage (see Fact A7), provided that leakage is independent of public coin a .

Theorem 5 (A security-preserving PRF on weak key). *Let n be a security parameter, and let $\mu = n^{-c}$ for constant $0 < c < 1$. Assume that the decisional LPN $_{\mu, n}$ problem is $(\alpha n, t, \epsilon)$ -hard for any constant $\alpha > 1$, then for any (efficiently computable) $d = \omega(1) \leq O(n)$ and any $q \leq n^{d/3}$ there exists a $(q, t - q\text{poly}(n), O(dq\epsilon))$ -randomized PRF on $(n^{1-\frac{c}{2}}, O(n^{1-c} \cdot \log n))$ -weak key*

$$\hat{\mathcal{F}} = \{\hat{F}_{k,a} : \{0, 1\}^n \rightarrow \{0, 1\}^n, k \in \{0, 1\}^{n^{1-\frac{c}{2}}}, a \in \{0, 1\}^{O(n^2)}\}_{n \in \mathbb{N}}$$

which are computable by a uniform family of polynomial-size depth- $O(d)$ circuits with unbounded-fan-in AND, OR and XOR gates.

Proof sketch. Following the proof of Theorem 4, we get a $(q, t - q\text{poly}(n), O(dq\epsilon))$ -randomized PRF $\mathcal{F} = \{F_{k,a}\}_{n \in \mathbb{N}}$ on weak keys (see (7)) with input length $d \log n$ and of depth $O(d)$. We define $\mathcal{F}' = \{F'_{(\mathbf{k}, \mathbf{h}), a} : \{0, 1\}^n \rightarrow \{0, 1\}^n, \mathbf{k} \in \{0, 1\}^{O(\kappa n^{1-\frac{c}{2}})}, \mathbf{h} \in \mathcal{H}^\kappa, a \in \{0, 1\}^{O(n^2)}\}_{n \in \mathbb{N}}$ where

$$F'_{(\mathbf{k}, \mathbf{h}), a}(x) \stackrel{\text{def}}{=} \bigoplus_{i=1}^{\kappa} F_{k_i, a}(h_i(x)), \mathbf{k} = (k_1, \dots, k_\kappa), \mathbf{h} = (h_1, \dots, h_\kappa).$$

Let $\delta_D(F_1, F_2) \stackrel{\text{def}}{=} |\Pr[\mathcal{D}^{F_1}(A) = 1] - \Pr[\mathcal{D}^{F_2}(A) = 1]|$. We have that for any oracle-aided distinguisher running in time $t - q\text{poly}(n)$ and making up to q queries, we have by a triangle inequality that

$$\begin{aligned} \delta_D(F'_{(\mathbf{K}, \mathbf{H}), A}, R) &\leq \delta_D(F'_{(\mathbf{K}, \mathbf{H}), A}, F_{R, \mathbf{H}}) + \delta_D(F_{R, \mathbf{H}}, R) \\ &\leq O(\kappa dq\epsilon) + n^{d(1-2\kappa)/3} \\ &= O(\kappa dq\epsilon) + 2^{-\omega(n^{1-c})} = O(\kappa dq\epsilon), \end{aligned}$$

where $F_{R, \mathbf{H}}$ is defined as per (8), the first term of the second inequality is due to a hybrid argument (replacing every $F_{K_i, A}$ with R_i one at a time), the second term of the second inequality follows from Lemma 5 with $\ell = d \log n$ and $q \leq n^{d/3}$, and the equalities follow by setting $\kappa = n^{1-c}$ to make the first term dominant. Therefore, $F'_{(\mathbf{k}, \mathbf{h}), a}$ is almost the PRF as desired except that it uses a long key (\mathbf{k}, \mathbf{h}) , which can be replaced with a pseudorandom one. That is, let $\tilde{F}_{k, a}(x) \stackrel{\text{def}}{=} F'_{(\mathbf{k}, \mathbf{h}), a}(x)$ and $(\mathbf{k}, \mathbf{h}) \stackrel{\text{def}}{=} F_{k, a}(1) \| F_{k, a}(2) \| \cdots \| F_{k, a}(O(\kappa))$, which adds only a layer of gates of depth $O(d)$. \square

5 An Alternative PRF with a Short Uniform Key

In this section, we introduce an alternative construction based on a variant of LPN (reducible from standard LPN) whose noise vector can be sampled in AC^0 (i.e., without using XOR gates). We state the end results in Theorem 6 that standard LPN with n -bit secret implies a low-depth PRF with key size $\Theta(n^{1-c} \log n)$. Concretely (and ideally), assume that computational LPN is $(q = 1.001n, t = 2^{n^{1-c}/3}, \epsilon = 2^{-n^{1-c}/12})$ -hard, and let $\lambda = \Theta(n^{1-c} \log n)$, then for any $\omega(1) = d = O(\lambda / \log^2 \lambda)$ there exists a parallelizable $(q' = \lambda^{\Theta(d)}, t' = 2^{\Theta(\lambda / \log \lambda)}, \epsilon' = 2^{-\Theta(\lambda / \log \lambda)})$ -randomized PRF computable in depth $O(d)$ with secret key length λ and public coin length $O(\lambda^{\frac{1+c}{1-c}})$.

5.1 Main Results and Roadmap

Theorem 6 (A PRF with a compact uniform key). *Let n be a security parameter, and let $\mu = n^{-c}$ for constant $0 < c < 1$. Assume that the computational $\text{LPN}_{\mu, n}$ problem is $(\alpha n, t, \epsilon)$ -hard for any constant $\alpha > 1$ and efficiently computable ϵ , then for any (efficiently computable) $d = \omega(1) \leq O(n)$ and any $q' \leq n^{d/3}$ there exists a $(q', \Theta(t \cdot \epsilon^2 n^{1-2c}), O(dq'n^2\epsilon))$ -randomized PRF on uniform key*

$$\tilde{\mathcal{F}} = \{\tilde{F}_{k, a} : \{0, 1\}^n \rightarrow \{0, 1\}^n, k \in \{0, 1\}^{\Theta(n^{1-c} \cdot \log n)}, a \in \{0, 1\}^{O(n^2)}\}_{n \in \mathbb{N}}$$

which are computable by a uniform family of polynomial-size depth- $O(d)$ circuits with unbounded-fan-in AND, OR and XOR gates.

We sketch the steps below to prove Theorem 6, where ‘C-’ and ‘D-’ stand for ‘computational’ and ‘decisional’ respectively.

1. Introduce distribution ψ_μ^m that can be sampled in AC^0 .
2. $((1 + \Theta(1))n, t, \epsilon)$ -hard C- $LPN_{\mu, n} \implies (\Theta(n), t - \text{poly}(n), 2\epsilon)$ -hard C- Ber_μ^{n+q} - $LPN_{\mu, n}$ (by Lemma 3).
3. $(\Theta(n), t, \epsilon)$ -hard C- Ber_μ^{n+q} - $LPN_{\mu, n} \implies (\Theta(n), t - \text{poly}(n), O(n^{3/2-c}\epsilon))$ -hard C- ψ_μ^{n+q} - $LPN_{\mu, n}$ (by Lemma 9).
4. $(\Theta(n), t, \epsilon)$ -hard C- ψ_μ^{n+q} - $LPN_{\mu, n} \implies (\Theta(n), \Omega(t(\epsilon/n)^2), 2\epsilon)$ -hard D- ψ_μ^{n+q} - $LPN_{\mu, n}$ (by Theorem 7).
5. $(\Theta(n), t, \epsilon)$ -hard D- ψ_μ^{n+q} - $LPN_{\mu, n} \implies (q, t - q \text{poly}(n), O(dq'\epsilon))$ -randomized PRF for any $d = \omega(1)$ and $q' \leq n^{d/3}$, where the PRF has key length $\Theta(n^{1-c} \log n)$ and can be computed by polynomial-size depth- $O(d)$ circuits with unbounded-fan-in AND, OR and XOR gates. This is stated as Theorem 8.

5.2 Distribution ψ_μ^m and the ψ_μ^{n+q} - $LPN_{\mu, n}$ Problem

We introduce a distribution ψ_μ^m that can be sampled in AC^0 and show that ψ_μ^{n+q} - $LPN_{\mu, n}$ is implied by Ber_μ^{n+q} - $LPN_{\mu, n}$ (and thus by standard LPN). Further, for $\mu = n^{-c}$ sampling ψ_μ^m needs $\Theta(mn^{-c} \log n)$ random bits, which asymptotically match the Shannon entropy of Ber_μ^m .

Algorithm 1. Sampling distribution ψ_μ^m in AC^0

Require: $2\mu m \log m$ random bits (assume WLOG that m is a power of 2)

Ensure: ψ_μ^m satisfies Lemma 6

- 1: Sample random $z_1, \dots, z_{2\mu m}$ of Hamming weight 1, i.e., for every $i \in [m]$ $z_i \stackrel{\$}{\leftarrow} \{z \in \{0, 1\}^m : |z| = 1\}$.
 {E.g., to sample z_1 with randomness $r_1 \dots r_{\log m}$, simply let each $(b_1 \dots b_{\log m})$ -th bit of z_1 to be $r_1^{b_1} \wedge \dots \wedge r_{\log m}^{b_{\log m}}$, where $r_j^{b_j} \stackrel{\text{def}}{=} r_j$ for $b_j = 0$ and $r_j^{b_j} \stackrel{\text{def}}{=} -r_j$ otherwise. Note that AC^0 allows NOT gates at the input level.}
 - 2: Output the bitwise-OR of the vectors $z_1, \dots, z_{2\mu m}$.
 {Note: we take a bitwise-OR (**not bitwise-XOR**) of the vectors.}
-

Lemma 6. *The distribution ψ_μ^m (sampled as per Algorithm 1) is $2^{-\Omega(\mu m \log(1/\mu))}$ -close to a convex combination of $\chi_{\mu m}^m, \chi_{\mu m+1}^m, \dots, \chi_{2\mu m}^m$.*

Proof. It is easy to see that ψ_μ^m is a convex combination of $\chi_1^m, \chi_2^m, \dots, \chi_{2\mu m}^m$ as conditioned on $|\psi_\mu^m| = i$ (for any i) ψ_μ^m hits every $y \in \{0, 1\}^m$ of Hamming weight $|y| = i$ with equal probability. Hence, it remains to show that those χ_j^m 's with Hamming weight $j < \mu m$ sum to a fraction less than $2^{-\mu m(\log(1/\mu)-2)}$, i.e.,

$$\begin{aligned} \Pr[|\psi_\mu^m| < \mu m] &= \sum_{y \in \{0, 1\}^m : |y| < \mu m} \Pr[\psi_\mu^m = y] \\ &< \mu^{2\mu m} \cdot 2^{m\mathbf{H}(\mu) - \frac{\log m}{2}} + O(1) \\ &< \mu^{2\mu m} \cdot 2^{\mu m(\log(1/\mu)+2)} + O(1) = 2^{\mu m(-\log(1/\mu)+2)} + O(1) \end{aligned}$$

where the first inequality is due to the partial sum of binomial coefficients (see Fact A5) and that for any fixed y with $|y| < \mu m$ $\psi_\mu^m = y$ happens only if the bit 1 of every z_i (see Algorithm 1) hits the 1's of y (each with probability less than μ independently) and the second inequality is Fact A1.

By definition of ψ_μ^{n+q} the sampled (s, e) has Hamming weight no greater than $2\mu(n+q)$ and the following lemma states that ψ_μ^{n+q} -LPN $_{\mu,n}$ is almost injective.

Lemma 7 (ψ_μ^{n+q} -LPN $_{\mu,n}$ is almost injective). *For $q = \Omega(n)$, define set $\mathcal{Y} \stackrel{\text{def}}{=} \{(s, e) \in \{0, 1\}^{n+q} : |(s, e)| \leq (n+q)/\log n\}$. Then, for every $(s, e) \in \mathcal{Y}$,*

$$\Pr_{a \leftarrow U_{qn}} [\exists (s', e') \in \mathcal{Y} : (s', e') \neq (s, e) \wedge as \oplus e = as' \oplus e'] = 2^{-\Omega(q)}.$$

Proof. Let $\mathcal{H} \stackrel{\text{def}}{=} \{h_a : \{0, 1\}^{n+q} \rightarrow \{0, 1\}^q, a \in \{0, 1\}^{qn}, h_a(s, e) \stackrel{\text{def}}{=} as \oplus e\}$ and it is not hard to see that \mathcal{H} is a family of universal hash functions. We have

$$\log |\mathcal{Y}| = \log \sum_{i=0}^{(n+q)/\log n} \binom{n+q}{i} = O((n+q) \log \log n / \log n) = o(q),$$

where the approximation is due to Fact A5 and the conclusion immediately follows from Lemma 8.

Lemma 8 (The injective hash lemma (e.g. [55])). *For any integers $l_1 \leq l_2, m$, let \mathcal{Y} be any set of size $|\mathcal{Y}| \leq 2^{l_1}$, and let $\mathcal{H} \stackrel{\text{def}}{=} \{h_a : \{0, 1\}^m \rightarrow \{0, 1\}^{l_2}, a \in \mathcal{A}, \mathcal{Y} \subseteq \{0, 1\}^m\}$ be a family of universal hash functions. Then, for every $y \in \mathcal{Y}$ we have*

$$\Pr_{a \leftarrow \mathcal{A}} [\exists y' \in \mathcal{Y} : y' \neq y \wedge h_a(y') = h_a(y)] \leq 2^{l_1-l_2}.$$

5.3 Computational Ber_μ^{n+q} -LPN $_{\mu,n} \rightarrow$ Computational ψ_μ^{n+q} -LPN $_{\mu,n}$

Lemma 9 non-trivially extends the well-known fact that the computational LPN implies the computational exact LPN, i.e., $(U_n, \chi_{\mu q}^q)$ -LPN $_{\mu,n}$.

Lemma 9. *Let $q = \Omega(n)$, $\mu = n^{-c}$ ($0 < c < 1$) and $\epsilon = 2^{-O(n^{1-c})}$. Assume that the computational Ber_μ^{n+q} -LPN $_{\mu,n}$ problem is (q, t, ϵ) -hard, then the computational ψ_μ^{n+q} -LPN $_{\mu,n}$ problem is $(q, t - \text{poly}(n+q), O(\mu(n+q)^{3/2}\epsilon))$ -hard.*

Proof. Let $m = n + q$ and write $\text{Adv}_D(X) \stackrel{\text{def}}{=} \Pr_{a \leftarrow U_{qn}, (s,e) \leftarrow X} [D(a, a \cdot s \oplus e) = (s, e)]$. Towards a contradiction we assume that there exists D such that $\text{Adv}_D(\psi_\mu^m) > \epsilon'$, and we assume WLOG that on input (a, z) D always outputs (s', e') with $|(s', e')| \leq 2\mu m$. That is, even if it fails to find any (s', e') satisfying $as' \oplus e' = z$ and $|(s', e')| \leq 2\mu m$ it just outputs a zero vector. Lemma 6 states that ψ_μ^m is $2^{-\Omega(\mu n \log(1/\mu))}$ -close to a convex combination of $\chi_{\mu m}^m, \chi_{\mu m+1}^m, \dots, \chi_{2\mu m}^m$, and thus there exists $j \in \{\mu m, \mu m + 1, \dots, 2\mu m\}$

such that $\text{Adv}_{\mathbb{D}}(\chi_j^m) > \epsilon' - 2^{-\Omega(n^{1-c} \log n)} > \epsilon'/2$, which further implies that $\text{Adv}_{\mathbb{D}}(\text{Ber}_{j/m}^m) = \Omega(\epsilon'/\sqrt{m})$ as $\text{Ber}_{j/m}^m$ is a convex combination of $\chi_0^m, \dots, \chi_m^m$, of which it hits χ_j^m with probability $\Omega(1/\sqrt{m})$ by Lemma 10. Next, we define \mathbb{D}' as in Algorithm 2.

Algorithm 2. A Ber_{μ}^m -LPN $_{\mu,n}$ solver \mathbb{D}'

Require: a random Ber_{μ}^m -LPN $_{\mu,n}$ instance $(a, z = a \cdot s \oplus e)$ as input

Ensure: a good chance to find out (s, e)

- 1: Sample $j^* \leftarrow^{\$} \{\mu m, \mu m + 1, \dots, 2\mu m\}$ as a guess about j .
 - 2: Compute $\mu' = j^*/m$.
 - 3: $(s_1, e_1) \leftarrow \text{Ber}_{\frac{\mu' - \mu}{1 - 2\mu}}^m$. {This makes $(a, z \oplus (as_1 \oplus e_1))$ a random $\text{Ber}_{\mu'}^m$ -LPN $_{\mu',n}$ sample by the piling-up lemma (see Fact A6)}
 - 4: $(s', e') \leftarrow \mathbb{D}(a, z \oplus (as_1 \oplus e_1))$.
 - 5: Output $(s' \oplus s_1, e' \oplus e_1)$. { \mathbb{D}' succeeds iff $(s' \oplus s_1, e' \oplus e_1) = (s, e)$ }
-

We denote \mathcal{E}_{suc} the event that \mathbb{D} succeeds in finding (s', e') such that $as' \oplus e' = z \oplus (as_1 \oplus e_1)$ and thus we have $a(s' \oplus s_1) \oplus (e' \oplus e_1) = z = as \oplus e$, where values are sampled as defined above. This however does not immediately imply $(s, e) = (s' \oplus s_1, e' \oplus e_1)$ unless conditioned on the event \mathcal{E}_{inj} that $h_a(s, e) \stackrel{\text{def}}{=} a \cdot s \oplus e$ is injective on input (s, e) .

$$\begin{aligned}
& \Pr[(s' \oplus s_1, e' \oplus e_1) = (s, e)] \\
& \stackrel{a \leftarrow U_{qn}, (s, e) \leftarrow \text{Ber}_{\mu}^m, (s_1, e_1) \leftarrow \text{Ber}_{\frac{\mu' - \mu}{1 - 2\mu}}^m, s' \leftarrow \mathbb{D}(a, y \oplus (as_1 \oplus e_1))}{\geq} \Pr[\mathcal{E}_{suc} \wedge \mathcal{E}_{inj}] \\
& \geq \Pr[\mathcal{E}_{suc}] - \Pr[\neg \mathcal{E}_{inj}] \\
& \geq \Pr[j^* = j] \cdot \text{Adv}_{\mathbb{D}}(\text{Ber}_{j/m}^m) - 2^{-\Omega(m/\log^2 n)} \\
& = \Omega(\epsilon'/\mu m^{3/2}),
\end{aligned}$$

where the bound on event $\neg \mathcal{E}_{inj}$ is given below. We reach a contradiction by setting $\epsilon' = \Omega(1) \cdot \mu m^{3/2} \epsilon$ for a large enough $\Omega(1)$ so that \mathbb{D}' solves Ber_{μ}^m -LPN $_{\mu,n}$ with probability greater than ϵ .

$$\begin{aligned}
& \Pr[\neg \mathcal{E}_{inj}] \\
& \leq \Pr[\neg \mathcal{E}_{inj} \wedge (s, e) \in \mathcal{Y} \wedge (s' \oplus s_1, e' \oplus e_1) \in \mathcal{Y}] \\
& \quad + \Pr[(s, e) \notin \mathcal{Y} \vee (s' \oplus s_1, e' \oplus e_1) \notin \mathcal{Y}] \\
& \leq 2^{-\Omega(m)} + \Pr[(s, e) \notin \mathcal{Y}] + \Pr[(s' \oplus s_1, e' \oplus e_1) \notin \mathcal{Y}] \\
& \leq 2^{-\Omega(m)} + \Pr_{(s, e) \leftarrow \text{Ber}_{\mu}^m}[|(s, e)| \geq m/\log n] + \Pr_{(s_1, e_1) \leftarrow \text{Ber}_{\frac{\mu' - \mu}{1 - 2\mu}}^m}[|(s_1, e_1)| \geq (\frac{1}{\log n} - 2\mu)m] \\
& = 2^{-\Omega(m/\log^2 n)},
\end{aligned}$$

where $\mathcal{Y} \stackrel{\text{def}}{=} \{(s, e) \in \{0, 1\}^m : |(s, e)| < m/\log n\}$, the second inequality is from Lemma 7, the third inequality is that $|(u \oplus w)| \geq \kappa$ implies $|w| \geq \kappa - |u|$ and by definition of D string (s', e') has Hamming weight no greater than $2\mu m$, and the last inequality is a typical Chernoff-Hoeffding bound.

Lemma 10. *For $0 < \mu' < 1/2$ and $m \in \mathbb{N}$, we have that*

$$\Pr \left[|\text{Ber}_{\mu'}^m| = \lceil \mu' m \rceil \right] = \Omega(1/\sqrt{m}).$$

5.4 C- ψ_{μ}^{n+q} -LPN $_{\mu,n} \rightarrow$ D- ψ_{μ}^{n+q} -LPN $_{\mu,n} \rightarrow \omega(1)$ -Depth PRFs

Next we show that the computational ψ_{μ}^{n+q} -LPN $_{\mu,n}$ problem implies its decisional counterpart. The theorem below is implicit in [5]⁹ and the case for ψ_{μ}^{n+q} -LPN $_{\mu,n}$ falls into a special case. Note that ψ_{μ}^{n+q} -LPN $_{\mu,n}$ is almost injective by Lemma 7, and thus its computational and decisional versions are equivalent in a sample-preserving manner. In fact, Theorem 7 holds even without the injective condition, albeit with looser bounds.

Theorem 7 (Sample preserving reduction [5]). *If the computational X-LPN $_{\mu,n}$ is (q, t, ϵ) -hard for any efficiently computable ϵ , and it satisfies the injective condition, i.e., for any $(s, e) \in \text{Supp}(X)$ it holds that*

$$\Pr_{a \leftarrow U_{qn}} [\exists (s', e') \in \text{Supp}(X) : (s', e') \neq (s, e) \wedge a \cdot s \oplus e = a \cdot s' \oplus e'] \leq 2^{-\Omega(n)}.$$

Then, the decisional X-LPN $_{\mu,n}$ is $(q, \Omega(t(\epsilon/n)^2), 2\epsilon)$ -hard.

Theorem 8 (Decisional ψ_{μ}^{n+q} -LPN $_{\mu,n} \rightarrow$ PRF). *Let n be a security parameter, and let $\mu = n^{-c}$ for any constant $0 < c < 1$. Assume that the decisional ψ_{μ}^{n+q} -LPN $_{\mu,n}$ problem is $(\delta n, t, \epsilon)$ -hard for any constant $\delta > 0$, then for any (efficiently computable) $d = \omega(1) \leq O(n)$ and any $q' \leq n^{d/3}$ there exists a $(q', t - q' \text{poly}(n), O(dq'\epsilon))$ -randomized PRF (on uniform key) with key length $\Theta(n^{1-c} \log n)$ and public coin size $O(n^2)$, which are computable by a uniform family of polynomial-size depth- $O(d)$ circuits with unbounded-fan-in AND, OR and XOR gates.*

Proof sketch. The proof is essentially the same as that of Theorem 5, replacing the Bernoulli randomness extractor with the ψ_{μ}^{n+q} sampler. That is, decisional ψ_{μ}^{n+q} -LPN $_{\mu,n}$ for $q = \Theta(n)$ implies a constant-depth polynomial-stretch randomized PRG on seed length $2\mu(n+q) \log(n+q) = \Theta(n^{1-c} \log n)$ and output length $\Theta(n)$, which in turn implies a nearly constant-depth randomized PRF, where the technique in Lemma 5 is also used to make the construction security preserving. □

⁹ Lemma 4.4 from the full version of [5] states a variant of Theorem 7 for uniformly random a and s , and arbitrary e . However, by checking its proof it actually only requires the matrix a to be uniform and independent of (s, e) .

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A Well-Known Facts, Lemmas and Inequalities

Fact A1. Let $\mathbf{H}(\mu) \stackrel{\text{def}}{=} \mu \log(1/\mu) + (1 - \mu) \log(1/(1 - \mu))$ be the binary entropy function. Then, for any $0 < \mu < 1/2$ it holds that

$$\mu \log(1/\mu) < \mathbf{H}(\mu) < \mu(\log(1/\mu) + 2).$$

Proof.

$$\begin{aligned} & \mu \log(1/\mu) \\ & < \left(\mathbf{H}(\mu) = \mu \log(1/\mu) + (1 - \mu) \log(1/(1 - \mu)) \right) \\ & = \mu \log(1/\mu) + (1 - \mu) \log\left(1 + \frac{\mu}{1 - \mu}\right) \\ & = \mu \log(1/\mu) + (1 - \mu) \frac{\ln\left(1 + \frac{\mu}{1 - \mu}\right)}{\ln 2} \\ & \leq \mu \log(1/\mu) + \frac{\mu}{\ln 2} < \mu(\log(1/\mu) + 2), \end{aligned}$$

where the first inequality is due to $(1 - \mu) \log(1/(1 - \mu)) > 0$, the second one follows from the elementary inequality $\ln(1 + x) \leq x$ for any $x > 0$, and the last inequality is simply $1 < 2 \ln 2$.

Lemma 11 (Chernoff bound). For any $n \in \mathbb{N}$, let X_1, \dots, X_n be independent random variables and let $\bar{X} = \sum_{i=1}^n X_i$, where $\Pr[0 \leq X_i \leq 1] = 1$ holds for every $1 \leq i \leq n$. Then, for any $\Delta_1 > 0$ and $0 < \Delta_2 < 1$,

$$\begin{aligned} \Pr[\bar{X} > (1 + \Delta_1) \cdot \mathbb{E}[\bar{X}]] &< \exp^{-\frac{\min(\Delta_1, \Delta_1^2)}{3} \mathbb{E}[\bar{X}]}, \\ \Pr[\bar{X} < (1 - \Delta_2) \cdot \mathbb{E}[\bar{X}]] &< \exp^{-\frac{\Delta_2^2}{2} \mathbb{E}[\bar{X}]}. \end{aligned}$$

Theorem 9 (The Hoeffding bound [30]). Let $q \in \mathbb{N}$, and let $\xi_1, \xi_2, \dots, \xi_q$ be independent random variables such that for each $1 \leq i \leq q$ it holds that $\Pr[a_i \leq \xi_i \leq b_i] = 1$. Then, for any $t > 0$ we have

$$\Pr \left[\left| \sum_{i=1}^q \xi_i - \mathbb{E} \left[\sum_{i=1}^q \xi_i \right] \right| \geq t \right] \leq 2 \exp^{-\frac{2t^2}{\sum_{i=1}^q (b_i - a_i)^2}}.$$

Fact A2. For any $\sigma \in \mathbb{N}^+$, the probability that a random $(n + \sigma) \times n$ Boolean matrix $M \sim U_{(n+\sigma) \times n}$ has full rank (i.e., rank n) is at least $1 - 2^{-\sigma+1}$.

Proof. Consider matrix M being sampled column by column, and denote \mathcal{E}_i to be the event that “column i is non-zero and neither is it any linear combination of the preceding columns (i.e., columns 1 to $i - 1$)”.

$$\begin{aligned} \Pr[M \text{ has full rank}] &= \Pr[\mathcal{E}_1] \cdot \Pr[\mathcal{E}_2 | \mathcal{E}_1] \cdots \Pr[\mathcal{E}_n | \mathcal{E}_{n-1}] \\ &= (1 - 2^{-(n+\sigma)}) \cdot (1 - 2^{-(n+\sigma)+1}) \cdots (1 - 2^{-(n+\sigma)+n-1}) \\ &> 2^{-(2^{-(n+\sigma)+1} + 2^{-(n+\sigma)+2} + \cdots + 2^{-(n+\sigma)+n})} \\ &> 2^{-2^{-\sigma+1}} \\ &> \exp^{-2^{-\sigma+1}} \\ &> 1 - 2^{-\sigma+1} \end{aligned}$$

where the first inequality is due to Fact A4 and the last follows from Fact A3.

Fact A3. For any $x > 0$ it holds that $\exp^{-x} > 1 - x$.

Fact A4. For any $0 < x < \frac{2-\sqrt{2}}{2}$ it holds that $1 - x > 2^{-(\frac{2+\sqrt{2}}{2})x} > 2^{-2x}$.

Fact A5 (A partial sum of binomial coefficients ([27], p. 492)). For any $0 < \mu < 1/2$, and any $m \in \mathbb{N}$

$$\sum_{i=0}^{m\mu} \binom{m}{i} = 2^{m\mathbf{H}(\mu) - \frac{\log m}{2} + O(1)}$$

where $\mathbf{H}(\mu) \stackrel{\text{def}}{=} \mu \log(1/\mu) + (1 - \mu) \log(1/(1 - \mu))$ is the binary entropy function.

Fact A6 (Piling-up Lemma). For any $0 < \mu \leq \mu' < 1/2$, $(\text{Ber}_\mu \oplus \text{Ber}_{\frac{\mu' - \mu}{1 - 2\mu}}) \sim \text{Ber}_{\mu'}$.

Fact A7 (Min-entropy source conditioned on leakage). Let X be any random variable over support \mathcal{X} with $\mathbf{H}_\infty(X) \geq l_1$, let $f : \mathcal{X} \rightarrow \{0, 1\}^{l_2}$ be any function. Then, for any $0 < \varepsilon < 1$, there exists a set $\mathcal{X}_1 \times \mathcal{Y}_1 \subseteq \mathcal{X} \times \{0, 1\}^{l_2}$ such that $\Pr[(X, f(X)) \in (\mathcal{X}_1 \times \mathcal{Y}_1)] \geq 1 - \varepsilon$ and for every $(x, y) \in (\mathcal{X}_1 \times \mathcal{Y}_1)$

$$\Pr[X = x \mid f(X) = y] \leq 2^{-(l_1 - l_2 - \log(1/\varepsilon))}.$$

B Lemmas and Proofs Omitted

Proof of Lemma 1. Recall that $\mathbf{H}(\mu) \stackrel{\text{def}}{=} \mu \log(1/\mu) + (1 - \mu) \log(1/(1 - \mu))$ equals to $\mathbf{H}_1(\text{Ber}_\mu)$. Parse Ber_μ^q as Boolean variables E_1, \dots, E_q , and for each $1 \leq i \leq q$ define

$$\xi_i \stackrel{\text{def}}{=} \begin{cases} 1, & \text{if } E_i = 1 \\ \frac{\log(\frac{1}{1-\mu})}{\log(\frac{1}{\mu})}, & \text{if } E_i = 0 \end{cases}$$

and thus we have that ξ_1, \dots, ξ_q are i.i.d. over $\{\frac{\log(1/(1-\mu))}{\log(1/\mu)}, 1\}$, each of expectation $\mathbf{H}(\mu)/\log(1/\mu)$.

$$\begin{aligned} & \Pr [\text{Ber}_\mu^q \in \mathcal{E}] \\ &= 1 - \Pr \left[\sum_{i=1}^q \xi_i > (1 + \Delta) \cdot \frac{q\mathbf{H}(\mu)}{\log(1/\mu)} \right] \\ &> 1 - \exp^{-\frac{\min(\Delta, \Delta^2)q\mathbf{H}(\mu)}{3\log(1/\mu)}} > 1 - \exp^{-\frac{\min(\Delta, \Delta^2)\mu q}{3}}, \end{aligned}$$

where the inequality follows from the Chernoff bound (see Lemma 11) and we recall $\mathbf{H}(\mu) > \mu \log(1/\mu)$ by Fact A1.

Proof of Lemma 3.

DECISIONAL $\text{LPN}_{\mu,n} \rightarrow$ DECISIONAL Ber_μ^{n+q} - $\text{LPN}_{\mu,n}$

Assume for contradiction there exists a distinguisher D that

$$\Pr_{A,S,E} [D(A, A \cdot S \oplus E) = 1] - \Pr_{A,U_{q-(n+2)}} [D(A, U_{q-(n+2)}) = 1] > 2\epsilon,$$

where $A \sim U_{(q-(n+2))n}$, $S \sim \text{Ber}_\mu^n$ and $E \sim \text{Ber}_\mu^{q-(n+2)}$. To complete the proof, we show that there exists another D' (of nearly the same complexity as D) that on input $(a', b) \in \{0, 1\}^{qn} \times \{0, 1\}^q$ that distinguishes $(A', A' \cdot X \oplus \text{Ber}_\mu^q)$ from (A', U_q) for $A' \sim U_{qn}$ and $X \sim U_n$ with advantage more than ϵ . We parse the $q \times n$ matrix a' and q -bit b as

$$a' = \begin{bmatrix} m \\ a \end{bmatrix}, \quad b = (b_m, b_a) \tag{9}$$

where m and a are $(n + 2) \times n$ and $(q - (n + 2)) \times n$ matrices respectively, $b_m \in \{0, 1\}^{n+2}$ and $b_a \in \{0, 1\}^{q-(n+2)}$. Algorithm D' does the following: it first checks whether m has full rank or not, and if not it outputs a random bit. Otherwise (i.e., m has full rank), D' outputs $D(a\bar{m}^{-1}, (a\bar{m}^{-1}) \cdot b_{\bar{m}} \oplus b_a)$, where \bar{m} is an $n \times n$ invertible submatrix of m and $b_{\bar{m}}$ is the corresponding¹⁰ substring of b_m . Now we give the lower bound of the advantage in distinguishing the two distributions. On the one hand, when $(a', b) \leftarrow (A', (A' \cdot X) \oplus \text{Ber}_\mu^q)$ and conditioned on that \bar{m} is invertible, we have that

$$\begin{aligned} \bar{m} \cdot x \oplus s &= b_{\bar{m}} \\ a \cdot x \oplus e &= b_a \end{aligned} \tag{10}$$

where $a \leftarrow U_{(q-(n+2))n}$, $x \leftarrow U_n$, $s \leftarrow \text{Ber}_\mu^n$, and $e \leftarrow \text{Ber}_\mu^{q-(n+2)}$, and it follows (by elimination of x) that $b_a = (a\bar{m}^{-1})s \oplus (a\bar{m}^{-1})b_{\bar{m}} \oplus e$, and thus $(a\bar{m}^{-1})b_{\bar{m}} \oplus b_a = (a\bar{m}^{-1})s \oplus e$. On the other hand, when $(a', b) \leftarrow (U_{qn}, U_q)$

¹⁰ E.g., if \bar{m} is the submatrix of m by keeping only the first n rows, then $b_{\bar{m}}$ is the n -bit prefix of b_m .

and conditioned on an invertible m it holds that $(a\bar{m}^{-1}, (a\bar{m}^{-1}) \cdot b_{\bar{m}} \oplus b_a)$ follows $(U_{(q-(n+2))n}, U_{q-(n+2)})$. Therefore, for $A \sim U_{(q-(n+2))n}$, $S \sim \text{Ber}_\mu^n$ and $E \sim \text{Ber}_\mu^{q-(n+2)}$ we have

$$\begin{aligned} & \Pr[\mathsf{D}'(U_{qn}, U_{qn} \cdot U_n \oplus \text{Ber}_\mu^q) = 1] - \Pr[\mathsf{D}'(U_{qn}, U_q) = 1] \\ & \geq \Pr[\mathcal{E}_f] \cdot \left(\Pr_{A,S,E}[\mathsf{D}(A, A \cdot S \oplus E) = 1] - \Pr_{A, U_{q-(1+\delta)n}}[\mathsf{D}(A, U_{q-(1+\delta)n}) = 1] \right) \\ & > (1 - 2^{-1})2\epsilon = \epsilon \end{aligned}$$

where \mathcal{E}_f denotes the event that $m \leftarrow U_{(n+2) \times n}$ has full rank whose lower bound probability is given in Fact A2.

COMPUTATIONAL LPN $_{\mu,n} \rightarrow$ COMPUTATIONAL Ber_μ^{n+q} -LPN $_{\mu,n}$

The reduction follows steps similar to that of the decisional version. Assume for contradiction there exists a distinguisher D that

$$\Pr_{A,S,E}[\mathsf{D}(A, A \cdot S \oplus E) = (S, E)] > 2\epsilon,$$

where $A \sim U_{(q-(n+2))n}$, $S \sim \text{Ber}_\mu^n$ and $E \sim \text{Ber}_\mu^{q-(n+2)}$, then there exists another D' that on input $(a', b = a'x \oplus e') \in \{0, 1\}^{qn} \times \{0, 1\}^q$ recovers (x, e') with probability more than ϵ . Similarly, D' parses (a', b) as in (9), checks if m has full rank and we define \bar{m} , $b_{\bar{m}}$ and \mathcal{E}_f same as the above reduction. Let $(s^*, e^*) \leftarrow \mathsf{D}(a\bar{m}^{-1}, (a\bar{m}^{-1}) \cdot b_{\bar{m}} \oplus b_a)$. As analyzed above, conditioned on \mathcal{E}_f we have $(a\bar{m}^{-1}) \cdot b_{\bar{m}} \oplus b_a = (a\bar{m}^{-1})s \oplus e$ where $(a\bar{m}^{-1}, s, e)$ follows distribution (A, S, E) defined above, and hence $(s^*, e^*) = (s, e)$ with probability more than 2ϵ . Once D' got s^* , it computes $x^* = \bar{m}^{-1} \cdot (b_{\bar{m}} \oplus s^*)$ (see (10)), $e'^* = a'x^* \oplus b$ and outputs (x^*, e'^*) .

$$\begin{aligned} & \Pr[\mathsf{D}'(A', A' \cdot X \oplus E') = (X, E')] \\ & \geq \Pr[\mathcal{E}_f] \cdot \Pr_{A,S,E}[\mathsf{D}(A, A \cdot S \oplus E) = (S, E)] \\ & > (1 - 2^{-1})2\epsilon = \epsilon \end{aligned}$$

where $A' \sim U_{qn}$, $X \sim U_n$ and $E' \sim \text{Ber}_\mu^q$. □

Proof of Lemma 5. To prove this indistinguishability result we use Patarin’s H-coefficient technique in its modern transcript-based incarnation [18, 48].

Without loss of generality the distinguisher D is deterministic and does not repeat queries. We refer to the case when the D ’s oracle is $F_{R,H}$ as the *real world* and to the case where the D ’s oracle is R as the *ideal world*.

D transcript consists of a sequence $(X_1, Y_1), \dots, (X_q, Y_q)$ of query-answer pairs to its oracle, plus (and following the “transcript stuffing” technique of [18]) the vector $\mathbf{H} = H_1, \dots, H_\kappa$ of hash functions, appended to the transcript after the distinguisher has made its last query; in the ideal world, \mathbf{H} consists of a “dummy” κ -tuple H_1, \dots, H_κ that can be sampled after the distinguisher’s last query, and is similarly appended to the transcript.

The probability space underlying the real world is $\Omega_{\text{real}} \stackrel{\text{def}}{=} \mathcal{H}^\kappa \times \mathcal{F}_{\ell \rightarrow n}^\kappa$ where $\mathcal{F}_{\ell \rightarrow n}$ is the set of all functions from ℓ bits to n bits, with uniform measure. The

probability space underlying the ideal world is $\Omega_{\text{ideal}} \stackrel{\text{def}}{=} \mathcal{H}^\kappa \times \mathcal{F}_{n \rightarrow n}$ where $\mathcal{F}_{n \rightarrow n}$ is the set of all functions from n bits to n bits, also with uniform measure.

We can identify elements of Ω_{real} and/or Ω_{ideal} as “oracles” for D to interact with. We write D^ω for the transcript obtained when D interacts with oracle ω , where $\omega \in \Omega_{\text{real}}$ in the real world and $\omega \in \Omega_{\text{ideal}}$ in the ideal world. Thus, the real-world transcripts are distributed according to $D^{W_{\text{real}}}$ where W_{real} is uniformly distributed over Ω_{real} , while the ideal-world transcripts are distributed according to $D^{W_{\text{ideal}}}$ where W_{ideal} is uniformly distributed over Ω_{ideal} .

A transcript τ is *attainable* if there exists some $\omega \in \Omega_{\text{ideal}}$ such that $D^\omega = \tau$. (Which transcripts are attainable depends on D , but we assume a fixed D). A transcript $\tau = ((X_1, Y_1), \dots, (X_q, Y_q), H_1, \dots, H_\kappa)$ is *bad* if there exists some $i \in [q]$ such that

$$H_j(X_i) \in \{H_j(X_1), \dots, H_j(X_{i-1})\}$$

for all $j \in \kappa$. We let T_{bad} be the set of bad attainable transcripts, T_{good} the set of non-bad attainable transcripts.

We will show that $\Pr[D^{W_{\text{real}}} = \tau] = \Pr[D^{W_{\text{ideal}}} = \tau]$ for all $\tau \in T_{\text{good}}$. In this case, by Patarin’s H-coefficient technique [18], D ’s distinguishing advantage is upper bounded by $\Pr[D^{W_{\text{ideal}}} \in T_{\text{bad}}]$. We commence by upper bounding the later quantity, and then move to the former claim.

Let $\mathcal{E}_{i,j}$, $(i, j) \in [q] \times [\kappa]$, be the event that

$$H_j(X_i) \in \{H_j(X_1), \dots, H_j(X_{i-1})\}$$

and let

$$\mathcal{E}_i = \mathcal{E}_{i,1} \wedge \dots \wedge \mathcal{E}_{i,\kappa}.$$

Since the values X_1, \dots, X_q and the hash functions H_1, \dots, H_κ are uniquely determined by any $\omega \in \Omega_{\text{ideal}}$ or $\omega \in \Omega_{\text{real}}$, we can write $\mathcal{E}_i(W_{\text{ideal}})$ (in the ideal world) or $\mathcal{E}_i(W_{\text{real}})$ (in the real world) to emphasize that \mathcal{E}_i is a deterministic predicate of the uniformly distributed oracle, in either world. Then

$$(D^{W_{\text{ideal}}} \in T_{\text{bad}}) \iff (\mathcal{E}_1(W_{\text{ideal}}) \vee \dots \vee \mathcal{E}_q(W_{\text{ideal}})). \tag{11}$$

Moreover,

$$\Pr[\mathcal{E}_{i,j}(W_{\text{ideal}})] \leq (i-1) \frac{1}{2^\ell} \leq \frac{q}{2^\ell}$$

since the hash functions H_1, \dots, H_κ are chosen independently of everything in the ideal world, and by the universality of \mathcal{H} , and

$$\Pr[\mathcal{E}_i(W_{\text{ideal}})] \leq \left(\frac{q}{2^\ell}\right)^\kappa$$

since the events $\mathcal{E}_{i,1}, \dots, \mathcal{E}_{i,\kappa}$ are independent in the ideal world; finally

$$\Pr[D^{W_{\text{ideal}}} \in T_{\text{bad}}] \leq q \left(\frac{q}{2^\ell}\right)^\kappa = \frac{q^{\kappa+1}}{2^{\ell\kappa}}$$

by (11) and by a union bound.

To complete the proof, we must show that $\Pr[D^{W_{\text{real}}} = \tau] = \Pr[D^{W_{\text{ideal}}} = \tau]$ for all $\tau \in T_{\text{good}}$. Clearly,

$$\Pr[D^{W_{\text{ideal}}} = \tau] = \frac{1}{2^{nq}} \cdot \frac{1}{|\mathcal{H}|^\kappa}$$

for all attainable τ . Moreover, if

$$\tau = ((x_1, y_1), \dots, (x_q, y_q), h_1, \dots, h_\kappa)$$

then it is easy to see that

$$\Pr[D^{W_{\text{real}}} = \tau \mid \mathbf{H}(W_{\text{real}}) = (h_1, \dots, h_\kappa)] = \frac{1}{2^{nq}}$$

by induction on the number of distinguisher queries, using $\tau \in T_{\text{good}}$. (We write $\mathbf{H}(W_{\text{real}})$ for the \mathbf{H} -coordinate of W_{real} .) Since

$$\Pr[\mathbf{H}(W_{\text{real}}) = (h_1, \dots, h_\kappa)] = \frac{1}{|\mathcal{H}|^\kappa}$$

this completes the proof. \square

Proof of Lemma 8.

$$\begin{aligned} & \Pr_{a \xleftarrow{\$} \mathcal{A}} [\exists y \in \mathcal{Y} : y' \neq y \wedge h_a(y') = h_a(y)] \\ & \leq \sum_{y' \in \mathcal{Y} \setminus \{y\}} \Pr_{a \xleftarrow{\$} \mathcal{A}} [h_a(y') = h_a(y)] \\ & \leq |\mathcal{Y}| \cdot 2^{-l_2} \leq 2^{-(l_2 - l_1)}, \end{aligned}$$

where the first inequality is a union bound and the second inequality follows by the universality of \mathcal{H} . \square

Proof of Lemma 10. Assume WLOG that $\mu' m$ is integer and use shorthand $p_l \stackrel{\text{def}}{=} \Pr[|\text{Ber}_{\mu'}^m| = l]$ and thus

$$p_{\mu' m} = \binom{m}{\mu' m} \mu^{\mu' m} (1 - \mu')^{m - \mu' m}$$

For $1 \leq i \leq \mu' m$, we have

$$\begin{aligned} p_{\mu' m - i} &= \binom{m}{\mu' m - i} \mu^{\mu' m - i} (1 - \mu')^{m - \mu' m + i} \\ &= \frac{m! \cdot \mu^{\mu' m} (1 - \mu')^{m - \mu' m}}{(\mu' m - i)! (m - \mu' m + i)!} \\ &= p_{\mu' m} \frac{(\mu' m - i + 1)(\mu' m - i + 2) \dots (\mu' m - i + i)}{(m - \mu' m + 1)(m - \mu' m + 2) \dots (m - \mu' m + i)} \cdot \left(\frac{1 - \mu'}{\mu'}\right)^i \\ &= p_{\mu' m} \frac{\left(1 - \frac{i-1}{\mu'}\right) \left(1 - \frac{i-2}{\mu'}\right) \dots \left(1 - \frac{0}{\mu'}\right)}{\left(1 + \frac{1}{m(1-\mu')}\right) \left(1 + \frac{2}{m(1-\mu')}\right) \dots \left(1 + \frac{i}{m(1-\mu')}\right)}. \end{aligned}$$

Similarly, for $1 \leq i \leq (1 - \mu')m$ we can show that

$$p_{\mu'm+i} = p_{\mu'm} \frac{(1 - \frac{0}{m(1-\mu')})(1 - \frac{1}{m(1-\mu')}) \dots (1 - \frac{i-1}{m(1-\mu')})}{(1 + \frac{1}{\mu'm})(1 + \frac{2}{\mu'm}) \dots (1 + \frac{i}{\mu'm})}.$$

Therefore, we have $p_{\mu'm} = \max\{p_i \mid 0 \leq i \leq m\}$ and thus complete the proof with the following

$$\begin{aligned} (1 + 2\sqrt{m}) \cdot p_{\mu'm} &\geq \sum_{j=\mu'm - \min\{\sqrt{m}, \mu'm\}}^{\mu'm + \sqrt{m}} p_j \\ &\geq 1 - \Pr[|\text{Ber}_{\mu'}^m| - \mu'm \geq \sqrt{m}] \\ &\geq 1 - 2\exp^{-2} = \Omega(1) \end{aligned}$$

where the last inequality is a Hoeffding bound. \square

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